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# Age-Related Sexual Dimorphism in **Temporal Discrimination and in** Adult-Onset Dystonia Suggests **GABAergic Mechanisms**

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Background: Adult-onset isolated focal dystonia (AOIFD) presenting in early adult life 068 is more frequent in men, whereas in middle age it is female predominant. Temporal dis-069 crimination, an endophenotype of adult-onset idiopathic isolated focal dystonia, shows 070 071 evidence of sexual dimorphism in healthy participants. 072

Objectives: We assessed the distinctive features of age-related sexual dimorphism of (i) 073 074 sex ratios in dystonia phenotypes and (ii) sexual dimorphism in temporal discrimination 075 in unaffected relatives of cervical dystonia patients. 076

077 Methods: We performed (i) a meta-regression analysis of the proportion of men in 078 published cohorts of phenotypes of adult-onset dystonia in relation to their mean age 079 of onset and (ii) an analysis of temporal discrimination thresholds in 220 unaffected 080 first-degree relatives (125 women) of cervical dystonia patients. 081

082 **Results:** In 53 studies of dystonia phenotypes, the proportion of men showed a highly 083 significant negative association with mean age of onset (p < 0.0001, pseudo- $R^2 = 59.6\%$ ). 084 with increasing female predominance from 40 years of age. Age of onset and phenotype 085 086 together explained 92.8% of the variance in proportion of men. Temporal discrimination 087 in relatives under the age of 35 years is faster in women than men but the age-related 088 rate of deterioration in women is twice that of men; after 45 years of age, men have faster 089 temporal discrimination than women. 090

091 **Conclusion:** Temporal discrimination in unaffected relatives of cervical dystonia patients 092 and sex ratios in adult-onset dystonia phenotypes show similar patterns of age-related 093 sexual dimorphism. Such age-related sexual dimorphism in temporal discrimination and 094 095 adult-onset focal dystonia may reflect common underlying mechanisms. Cerebral GABA 096 levels have been reported to show similar age-related sexual dimorphism in healthy 097 participants and may be the mechanism underlying the observed age-related sexual 098 dimorphism in temporal discrimination and the sex ratios in AOIFD. 099 100

Keywords: adult-onset isolated focal dystonia, sex ratio, temporal discrimination, sexual dimorphism, penetrance

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### 101 INTRODUCTION

Twenty years ago, a paper from the late David Marsden's research
group, reporting on the sex ratios of the phenotypes of adult-onset
dystonia, concluded by stating: "Our study confirms a clear but
mild preponderance of females with various types of craniocervical
dystonia and of males with writer's cramp. Why this is so remains
to be discovered" (1).

Dystonia is a movement disorder, characterized by "sustained 109 or intermittent muscle contractions causing abnormal, often 110 repetitive, movements, postures, or both" (2). The most common 111 form of dystonia, adult-onset isolated focal dystonia (AOIFD) 112 is inherited in an autosomal dominant manner with a reduced 113 penetrance of 12-15% (3, 4); phenotypes include cervical dysto-114 nia, blepharospasm, focal hand dystonia, spasmodic dysphonia, 115 oromandibular dystonia, and task-specific dystonia. Evidence 116 from studies of affected sib-pairs (5) and multiplex families (6, 117 7) indicates that the same presumed genetic mutation(s) may 118 cause different phenotypes. There is an unexplained male pre-119 dominance in focal hand dystonia and musician's dystonia with a 120 female excess in the craniocervical phenotypes. 121

The temporal discrimination threshold (TDT) is the shortest 122 interval at which two sequential stimuli appear to the observer 123 to be asynchronous; normally, the TDT is 30-50 ms depending 124 on the gender and age (8). Abnormal TDTs are found in 97% of 125 cervical dystonia patients (specificity: 98–100%) and somewhat 126 less frequently in other AOIFD phenotypes (8–10). Abnormal 127 TDTs show autosomal dominant transmission in families of spo-128 radic and familial cervical dystonia patients (8, 11). In unaffected 129 first-degree female relatives, an abnormal TDT is fully penetrant 130 by 48 years of age; in male relatives, there is 40% penetrance after 131 25 years of age (12). Abnormal TDTs in unaffected first-degree 132 relatives of patients with AOIFD are associated with putaminal 133 hypertrophy (8) and reduced putaminal fMRI activity during 134 a temporal discrimination task (12). We have proposed that an 135 abnormal TDT is a mediational endophenotype of AOIFD and 136 that unaffected first-degree relatives (of AOIFD patients) with 137 abnormal TDTs are unaffected gene carriers (13, 14). 138

In a group of healthy control participants, we observed that young women had faster temporal discrimination than similarly aged men (12). Further analysis indicated both sex-related and age-related effects on the TDT. TDTs worsened with age at a faster rate in women than men, so that in middle age the initial female advantage was lost; women after 40 years were significantly slower in a temporal discrimination task than men (15).

Sexual dimorphism, a phenotypic difference between males 146 and females of the same species, in temporal discrimination is age 147 dependent; the initial advantage of the young adult woman is lost 148 and reversed in middle age. In this paper, we further examine this 149 sexual dimorphism in temporal discrimination in a large group of 150 unaffected first-degree relatives of cervical dystonia patients. We 151 also examine by meta-analysis the relationship between mean age 152 of onset and sex ratios in published cohorts of AOIFD phenotypes. 153 154

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Abbreviations: AOIFD, adult-onset isolated focal dystonia; TDT, temporal dis-crimination threshold.

We postulate that the age-related sexual dimorphism in temporal 158 discrimination explains the relationship between and the mean 159 age of onset and sex ratios of AOIFD phenotypes in both published 160 clinical cohorts and in our cervical dystonia patients. We hypoth-161 esize that, in individuals with a genetic susceptibility to AOIFD, 162 age- and sex-related mechanisms relating to the efficiency of tem-163 poral discrimination explain the observed sex ratios in the various 164 AOIFD phenotypes. One further corollary hypothesis, discussed 165 but not examined in this paper, is that sexual dimorphism in the 166 speed of temporal discrimination reflects sexual dimorphism in 167 physiological age-related decline in GABAergic inhibition. 168

### PARTICIPANTS AND METHODS

### Study A: Age of Onset, Phenotypes, and Sex Ratios in AOIFD Study A Population

We searched, in January 2015, PubMed and MEDLINE for pub-176 lications using the search terms "adult-onset dystonia," "focal 177 dystonia," "blepharospasm," "cervical dystonia," focal hand 178 dystonia," "writer's cramp," "spasmodic dysphonia," "laryn-179 geal dystonia," oromandibular dystonia," "Meige syndrome," 180 "muscician's dystonia," "sex ratio," "dystonia," "epidemiology," 181 "incidence," and "prevalence." Our inclusion criteria were 182 publications that reported the numbers of patients, sex ratios 183 (proportion of men), and mean age of onset in clearly defined 184 AOIFD phenotypes from a clinic or study sample population. 185 Three authors (Ines M. Beiser, Michael Hutchinson, and Seán 186 O'Riordan) screened full texts to identify study eligibility. 187 We included only studies that were published in English; the 188 references of all eligible studies were searched to ensure that 189 no study was missed. Exclusion criteria were studies which did 190 not separate secondary dystonias or dystonia-plus syndromes 191 and repeated studies from the same geographical region (in 192 which case the largest or most recent study was used). A table of 193 included studies, listed by phenotype and references, is given in 194 Table S1 in Supplementary Material. 195

### Statistical Analysis

198 A mixed effects meta-regression model was fitted to the logit of 199 the sex ratio (proportion of cases that were male to total number 200 of cases) with (i) mean age of onset and (ii) mean age of onset plus 201 phenotype being included as moderators in the model. The model 202 was fitted in R version 3.1.1 using the package metafor version 203 1.9.3 (16). In reporting, the model's estimates of the effect sizes and SEs, together with pseudo- $R^2$ , defined as  $\left(\tau_{RE}^2 - \tau_{ME}^2\right) / \tau_{RE}^2$ 204 205 are presented (17). 206

## Study B: Age- and Sex-Related Effects on Temporal Discrimination in Relatives of Cervical Dystonia Patients Participants

Two hundred twenty unaffected first-degree relatives of cervical dystonia patients between the ages of 18 and 65 years (125 213 women, mean age 40.2 years; 95 men, mean age 39.7 years) were 214

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215 recruited by initial contact with cervical dystonia patients and subsequently gave full informed consent. A full medical history 216 was taken, and the relatives were assessed for any evidence of a 217 218 neurological disorder. Exclusion criteria were a history of neurological disease, including dystonia, tremor, neuropathy, visual 219 or cognitive impairment, a history of cerebral, and cervical or 220 brachial plexus injury. The TDT results from 175 healthy control 221 participants (88 women, mean age 40.5 years; 87 men, mean 2.2.2 age 21.5 years) from a previous study were included as a control 223 group (15). 224

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### 226 Methods: Sensory Testing:

### 227 Visual and Tactile TDT Testing

Methods: sensory testing: visual and tactile TDT testing was 228 229 performed in a single session; in a soundproof, darkened room 230 as described previously. Visual stimuli (two flashing LED lights) were positioned on a table, 7° in the participant's periphery. 231 Tactile stimuli (non-painful electrical impulses to the index and 232 middle fingers) were presented using square-wave stimulators 233 (Lafayette Instruments Europe, United Kingdom) and rectangu-234 lar cloth electrodes (Item # TD-141C1, Discount Disposables, St. 235 Albans, VT, USA). The stimulus current was manually increased 236 (in 0.1 mA steps) until the participant could reliably detect the 237 stimuli. Visual or tactile stimuli, 5 ms in duration, were presented 238 at 5-s intervals. The stimuli were initially synchronous and separa-239 tion between pairs of stimuli was introduced in 5 ms steps. When 240 the participant reported stimuli to be asynchronous on three 241 consecutive occasions, the first of these was taken as the TDT. 242 Visual and tactile testing was repeated four times on each side of 243 244 the body (a total of 16 runs) in a random order and the median (milliseconds) of the four trials was used to account for a practice 245 effect. Means of the median visual, tactile, and combined values 246 were calculated (TDT). Testing was carried out by the research 247 registrars according to a standardized protocol. 248

### 250 Statistical Analysis

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To investigate the effect of age and sex on temporal discrimina-251 tion, regression analyses were performed. The combined TDTs 2.52 for men and for women were submitted to regression analyses 253 with age as the continuous variable. The F values, R<sup>2</sup> values, and 254 corresponding *p* values are reported along with 95% confidence 255 intervals, *t*-values, and *p* values for the intercept; and  $\beta$  value for 256 257 the linear fit. To compare the intercept and  $\beta$  values between men 258 and women, a regression analysis was performed on the TDT data with variables age, sex (men = 0, women = 1) and age  $\times$  sex 259 (resulting in 0 s for men and the continuous variable of age for 260 women). The sex variable tests for differences in the intercept 261 values between relatives and women. The age  $\times$  sex variable tests 262 for differences in the β values between men and women. To inves-263 tigate intercept and  $\beta$  values differences between relatives and 264 265 control participants, a regression analysis was performed on the TDT data with variables age, group (relatives = 0, controls = 1) 266 and age  $\times$  sex (resulting in 0 s for relatives and the continuous 267 variable of age for controls). The group variable tests for differ-268 ences in the intercept values between relatives and controls. The 269 age  $\times$  group variable tests for differences in the  $\beta$  values between 270 relatives and controls. Approval for this project was obtained 271

from the Ethics and Medical Research Committee, St. Vincent's 272 University Hospital. 273

# Study C: Cervical Dystonia: Difference in Mean Age of Onset by Sex

### **Statistical Analysis**

Meta-analysis of the mean age of onset by sex in our cohort and<br/>three other studies (18–20), which reported mean age of onset<br/>of cervical dystonia for both sexes, was performed and the mean<br/>difference in age of onset between men and women presenting<br/>with cervical dystonia was determined. A random effects meta-<br/>analysis was performed using Review Manager version 5.3,<br/>Cochrane database tool (http://tech.cochrane.org/revman).284<br/>285<br/>286<br/>286<br/>287<br/>288<br/>289<br/>290<br/>290<br/>291

# RESULTS

### Study A: Age of Onset, Phenotypes, and Sex Ratios in AOIFD Analysis Set

A total of 78 papers fulfilling the search criteria were found. Of 298 these, 24 studies listed both the mean age of onset and the sex 299 ratio for the AOIFD phenotypes of interest; there were 54 papers 300 which did not provide adequate information for analysis or were 301 repeated studies from the same site/region. The 24 included stud-302 ies generated a total of 53 reports of mean age of onset and sex 303 ratio for the each of the five phenotypes: cervical dystonia (15 304 reports), focal hand dystonia (12 reports), musician's dystonia (3 305 reports), larvngeal dystonia (9 reports), and blepharospasm (14 306 reports). A list of all analyzed studies and phenotypes is available 307 in Table S1 in Supplementary Material. 308

### Meta-Regression Analysis of Relationship Between Sex Ratio and Mean Age of Onset

For all phenotypes (53 reports of phenotypes from 24 studies), 312 using a mixed effects meta-regression analysis, a statistically 313 significant association between the sex ratio (proportion men) 314 and mean age of onset was identified; the coefficient of age of 315 onset in a model for logit (proportion men) was -0.0709, (95%) 316 CI -0.0900 to -0.0518) p < 0.0001, (**Table 1**). The pseudo- $R^2$ 317 value for this association was 59.63%. This indicates evidence 318 of a linear decreasing male sex ratio with increasing age at 319 symptom onset and that almost 60% of the variance in sex 320 ratio is accounted for by age at onset. This effect is illustrated 321 in the plot of the proportion of men versus mean age of onset 322 in Figure 1. 323

### Meta-Regression Analysis of Relationship Between Sex Ratio and Phenotype at Presentation

Using the same dataset of 53 reports in a mixed effects metaregression analysis, a statistically significant association between 328

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	Coefficient	SE	Z value	p Value	Confidence intervals		
(A) META-REGRESSION ANALYSIS OF SEX RATIO (PROPORTION OF MEN) IN RELATION TO MEAN AGE OF ONSET OF ADULT-ONSET DYSTONIA							
ntercept	2.8780	0.4487	6.4141	< 0.0001	1.9986	3.7574	
Age of onset	-0.0709	0.0098	-7.2655	<0.0001	-0.0900	-0.0518	
B) META-REGRESSION	ANALYSIS OF SEX RATI	O (PROPORTION O	F MEN) IN RELATION	TO DYSTONIA PHENO	TYPE WITH BLEPHAR	OSPASM AS	
Intercept	-0.8387	0.1101	-7.6183	< 0.0001	-1.0545	-0.6229	
Cervical dystonia	0.351	0.1451	2.42	0.0155	0.067	0.6353	
Focal hand dystonia	1.17	0.17	6.93	< 0.0001	0.84	1.50	
Laryngeal dystonia	0.047	0.19	0.243	0.808	-0.33	0.42	
Musician's dystonia	2.30	0.27	8.50	<0.0001	1.77	2.83	
(C) META-REGRESSION	ANALYSIS OF SEX RATI	O (PROPORTION O	F MEN) IN RELATION	TO BOTH MEAN AGE	OF ONSET OF ADULT-	ONSET	
Intercept	4.7640	1.1084	4.2982	< 0.0001	2,5916	6.9364	
Age of onset	-0.1002	0.0197	-5.0840	< 0.0001	0.1388	0.0616	
Cervical dystonia	-1.0618	0.3013	-3.5239	0.0004	1.6523	0.4712	
Focal hand dystonia	-0.4995	0.3556	-1.4046	0.1601	-1.1966	0.1975	
Laryngeal dystonia	-0.8752	0.2386	-3.6685	0.0002	1.3429	0.4076	
Musician's dystonia	0.0306	0.5006	0.0611	0.9513	-0.9505	1.0117	

TABLE 1 | Meta-regression analysis of relationship between mean age of onset and sex ratio in phenotypes of adult-onset dystonia

348 Meta-regression analysis of the two independent variables of interest: (a) mean age of onset of dystonia, (b) adult-onset dystonia phenotype, and (c) combined variables [(a) and 349 (b)] with the sex ratio (proportion of men) as the dependent variable. Coefficients, SEs, p values, and 95% confidence intervals are shown; p values <0.05 are considered to be statistically significant.

the sex ratio (proportion of men) and phenotype was identified; the coefficient of phenotypes in the model for logit (proportion men) are presented in **Table 1**. The pseudo- $R^2$  value for this association was 81%.

# <sup>356</sup> Meta-Regression Analysis of Relationship Between <sup>357</sup> Sex Ratio and Both Mean Age of Onset and

# <sup>336</sup> Phenotype at Presentation

Using the same dataset of 53 reports in a mixed effect meta-360 regression analysis, a statistically significant association between 361 the sex ratio (proportion of men) and both age at onset and 362 phenotype was identified (**Table 1**). The pseudo- $R^2$  value for this 363 association was 92.83%. This indicates that the almost 93% of the 364 variance in sex ratio (proportion men) in adult-onset dystonia 365 is accounted for by the combination of mean age at onset and 366 phenotype. A model, which included interactions between phe-367 notype and age, was also fitted but these were not found to be 368 statistically significant. 369

# Study B: Age- and Sex-Related Effects on Temporal Discrimination in Unaffected Polatives

# **373 Unaffected Relatives**

The linear regression fits of TDT as a function of age for women 374 (red) and men (blue) in 220 unaffected first-degree relatives (solid 375 lines) and 175 control participants (broken lines) are illustrated 376 in Figure 2. The regression for female relatives showed that age 377 explained a significant amount of the variance in the TDT values 378 379  $[F(1,124) = 45.758, p < 0.001, R^2 = 0.271]$ , with a significant intercept 12.7 ms,  $\{t(124) = 2.08, p < 0.05; 95\%$  CI [8.254, 24.814]}, and 380 a significant  $\beta = 0.985$  ms, {t(124) = 6.76, p < 0.001; 95% CI [0.697, 381 1.273]} (Table S2 in Supplementary Material). In male relatives, 382 there was a significant relationship between age and TDT values 383  $[F(1,94) = 8.9, p < 0.005, R^2 = 0.087]$  demonstrating a significant 384 intercept 31.09 ms,  $\{t(94) = 5.065, p < 0.001; 95\% \text{ CI} [18.9, 43.283]\},\$ 385

408 and a significant  $\beta = 0.437 \text{ ms}, \{t(94) = 2.984, p < 0.005; 95\% \text{ CI} [0.22], t = 0.005; 0.005$ 409 0.738]}. Comparison of regression analysis in relatives between 410 women and men revealed a significantly lower intercept in women 411 than men  $-18.382 \text{ ms} \{t(218) = -2.11, p < 0.05; 95\% \text{ CI} [-35.49]$ 412 -1.273] and significant  $\beta$  value for the interaction of sex and age 413  $0.548 \text{ ms} \{t(218) = 2.646, p < 0.01; 95\% \text{ CI} [0.14, 0.956]\}, \text{ with a}$ 414 1 ms increase in the TDT every year in women; in men an increase 415 by 0.5 ms each year. The regression line for men and the regression 416 line for women intersected at approximately 31 years of age. As 417 with the unaffected relatives, the regression analysis between men 418 and women in the 175 control participants showed an effect of sex 419 and an interaction of sex and age [see Table S2 in Supplementary 420 Material and Ref. (15)]. Comparison of regression fits of control 421 participants and relatives resulted in a significant difference in the  $\beta$ 422 value for the interaction of age and group  $-0.343 \{t(392) = -2.293, t(392) \}$ 423 p < 0.05; 95% CI [-0.594, -0.049]}, with no significant effect of 424 group (Table S2 in Supplementary Material). While both unaffected 425 relatives and control participants at age 20 years have similar TDTs, 426 with increasing age the TDT worsens by 0.34 ms/year in relatives 427 compared to healthy control participants. 428

## Study C: Cervical Dystonia: Differential Mean Age of Onset by Sex

Meta-analysis of the mean age of onset by sex in our cervi-432 cal dystonia cohort in 278 patients (188 women) and 3 other 433 studies (18–20), which reported mean age of onset of cervical 434 dystonia for both sexes is shown in **Figure 3**. In each of the four 435 cohorts (total: 303 men and 506 women), men had an earlier 436 mean age of onset than women; in the meta-analysis, the mean 437 age of onset of cervical dystonia was significantly earlier by 438 4.3 years (95% CI; 1.8–6.8) in men than women (p = 0.0008). 439 Thus, despite the overall preponderance of women in cervical 440 dystonia, men have a significantly earlier mean age of onset 441 than women. 442

Proportion of men to total number of cases 1.0 0.8 0.6 0.4 0.2 0.0 age of onset of adult onset dystonia (years) FIGURE 1 | Meta-regression analysis of the relationship between sex ratio (proportion of men) and mean age of onset in phenotypes of adult-onset 

**dystonia**. Fifty-three reports of mean age of onset and sex ratio of five adult-onset dystonia phenotypes (listed in Table S1 in Supplementary Material). In each filled circle, the diameter of the circle is proportional to the square root of the number of individuals in that study. Circle color indicates musician's dystonia (pink), focal hand dystonia (blue), cervical dystonia (green), laryngeal dystonia (turquoise), and blepharospasm (red). A linear decreasing proportion of men (reducing male:female ratio) with increasing mean age of onset is statistically significant (p < 0.0001). The pseudo- $R^2$  value for this association is 59.63%. The solid line is the fitted meta-regression back transformed from the logit scale. The dashed lines are the 95% confidence intervals for the mean association back transformed to the original scale.



FIGURE 2 | Sexual dimorphism in temporal discrimination in unaffected first-degree relatives of cervical dystonia patients. The effect of age and gender on the temporal discrimination threshold (in milliseconds) in 220 unaffected first-degree relatives of patients with cervical dystonia. Each open circle represents an individual relative's temporal discrimination threshold (men: blue; women: red). The solid lines represent the regression fit of the data for men (blue) and women (red) of temporal discrimination thresholds with age. The broken red and blue lines indicate the regression fit of the temporal discrimination threshold for 175 healthy control participants [from Ref. (15)]. The sexually dimorphic effects of age on the temporal discrimination thresholds in men and women are statistically significant for both the unaffected first-degree relatives and the healthy control participants.

# DISCUSSION

# Age-Related Sexual Dimorphism in Disease Penetrance in AOIFD

The male:female sex ratio (proportion of men) in AOIFD decreases with increasing mean age at onset; this association is highly significant and mean age of onset accounts for almost 60% of the variance in the proportion of men. AOIFD with onset below the age of 40 years affects predominantly men in focal hand dystonia and musician's dystonia. After 40-45 years of age, the AOIFD phenotypes predominantly affect women and, importantly, with increasing age there is a steady linear decrease in the proportion of men affected (Figure 1). The cervical dystonia phenotype is of particular interest because, although overall it affects women 1.5-2.0 times more frequently than men, men have a mean age of onset approximately 4 years earlier than women (Figure 3). 

# Age-Related Sexual Dimorphism in Temporal Discrimination in Unaffected Relatives

In both unaffected first-degree relatives and healthy control 551 participants (15), temporal discrimination shows a significant 552 effect of sex, age, and an interaction of sex and age. Female 553 relatives and female control participants under the age of 554



571 40 years have faster temporal discrimination than age-matched 572 male relatives and controls. Overall temporal discrimination 573 worsens as a function of age but this is more pronounced 574 in women than men; the TDT in women deteriorates two 575 times faster than in men. Sexual dimorphism of temporal 576 discrimination is thus seen both in unaffected relatives 577 (Figure 2) and in healthy participants (15). However, the 578 effect of age on the TDT of unaffected relatives was more 579 marked than in healthy participants, worsening by 0.34 ms/ 580 year more in relatives than controls. We postulate that this 581 increased worsening with age in unaffected relatives indicates 582 an effect of non-manifesting gene carriage on the TDT endo-583 phenotype; approximately half of these unaffected first-degree 584 relatives are non-manifesting carriers of susceptibility genes 585 for adult-onset dystonia (13, 14). 586

# <sup>587</sup> Age-Related Sexual Dimorphism in <sup>588</sup> Ambient GABA

GABA levels, measured by magnetic resonance spectroscopy (MRS), in the frontal lobes in healthy participants show agerelated sexual dimorphism (21). Under 40 years of age women have higher GABA levels than men. However, with increasing age, women show more marked progressive reduction in frontal lobe GABA than men, so that after 40 years of age men have increasingly higher frontal lobe GABA than women.

Deficient inhibitory GABAergic mechanisms are considered 597 to be important in the pathogenesis of dystonia. Cortical and 598 subcortical GABA levels are reduced contralateral to the affected 599 hand in patients with focal hand dystonia (22). Alterations in 600 inhibitory circuits in dystonia have been noted at spinal cord, 601 brainstem, and cortical levels (23). Transcranial magnetic stimu-602 lation studies have shown impaired intra-cortical inhibition in 603 both hemispheres in focal hand dystonia, even though symptoms 604 were present on only one side of the body (24). Deficient inhibi-605 tion may be responsible for excessive and overflow movements 606 seen in AOIFD (24, 25). 607

# <sup>609</sup> Age, Cognition, and GABA

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Normal aging is associated with a decline in aspects of cognitive
function, including slowing in performance of visual orientation
discrimination (26), visual face matching (27), three-dimensional
shape determination (28), and motion direction detection (29).

628 Even in young healthy individuals, variation in cortical GABA 629 levels has been found to relate to a number of psychophysical 630 measures, including tactile discrimination thresholds (30) and 631 luminance grating orientation (31), with poorer performance 632 associated with regionally reduced GABA levels. The ability to 633 respond to a visual target speedily in the presence of a visual dis-634 traction has been linked to frontal lobe GABA measured by MRS 635 (32). Reduced GABAergic inhibition, either age-related or due to 636 physiological variation, results in blunting of stimulus-response 637 tuning specificity and task performance. In animals, degradation 638 in these visual discriminatory abilities, associated with a lower 639 signal to noise ratio in the striate cortex, has been linked to a 640 reduction (by 45-60%) of GABA immuno-reactive neurones in 641 the older animal (33, 34). Age-related reduction in GABA has 642 been noted in the inferior colliculus (35), and the hippocampus 643 in rats (36).

# Linking Clinical and Laboratory Observations

We postulate that the age-related sexual dimorphism in (a) dis-648 ease penetrance in AOIFD, (b) temporal discrimination, and (c) 649 ambient cerebral GABA are linked and illuminate pathogenic 650 mechanisms in AOIFD. The underlying hypothesis is that the 651 efficiency of temporal discrimination is a measure of effective 652 temporal tuning in the basal ganglia/superior colliculus, and 653 relates to GABAergic inhibitory mechanisms. Experimental 654 work in animals indicates that effective GABAA and GABAB 655 inhibitory mechanisms are necessary for the sharp onset and 656 offset of responses to visual stimuli in the superficial laminae 657 of the superior colliculus (37-39). While GABA may be meas-658 ured in large cortical structures by MRS, there is no method, 659 as yet, to measure ambient or synaptic GABA in a structure 660 as small as the superior colliculus in man in vivo. Thus, the 661 hypothesis that faster temporal discrimination reflects more 662 effective GABAergic inhibition in man remains likely, but 663 unproven. The clinical relevance to adult-onset focal dystonia is 664 that age-related sexual dimorphism in ambient GABA (causing 665 the observed sexual dimorphism in temporal discrimination), 666 results in the age-related sexual dimorphism in AOIFD disease 667 penetrance. 668

Arising from these observations, we hypothesize that, in a 669 person with a genetic susceptibility to AOIFD, penetrance varies 670

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according to gender and age (Figure 4): (i) under 40 years of age, 701 women have faster temporal discrimination than men, have more 702 ambient GABA and are resistant to disease penetrance; thus, 703 the observed male predominance in focal hand dystonia and 704

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musician's dystonia. (ii) The increased age-related rate of dete- 728 rioration in temporal discrimination in women reflects faster 729 reducing ambient GABA levels in women than men. After 730 40-45 years, women have increasingly worse temporal discrimi- 731 nation and increasingly lower ambient GABA than men. Older 732 women are thus more susceptible to disease penetrance and thus 733 the observed progressively increasing age-related female pre-734 dominance in adult-onset dystonia phenotypes presenting after 735 the age of 40 years (Figure 4). 736

# AUTHOR CONTRIBUTIONS

740 Conception (MH); design (MH, RR, and SO); organization (JB, 741 LW, IB, and EM); statistical analysis (CW, JB, and LW); execution 742 (IB, LW, EM, JB, DH, TL, HM, FM, and RW); review and critique 743 (RR, MH, and JB); manuscript preparation: writing the first draft 744 (MH); review and critique (all authors). 745

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# SUPPLEMENTARY MATERIAL

- The Supplementary Material for this article can be found online at 760 http://journal.frontiersin.org/article/10.3389/fneur.2015.00258
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